

- 1) Introduction to clustering
- 2) Partitioning Methods
 - K-Means
 - K-Medoid
 - Choice of parameters: Initialization, Silhouette coefficient
- 3) Expectation Maximization: a statistical approach
- 4) Density-based Methods: DBSCAN
- 5) Hierarchical Methods
 - Agglomerative and Divisive Hierarchical Clustering
 - Density-based hierarchical clustering: OPTICS
- 6) Evaluation of Clustering Results
- 7) Further Clustering Topics
 - Ensemble Clustering
 - Discussion: an alternative view on DBSCAN

Statistical approach for finding maximum likelihood estimates of parameters in probabilistic models

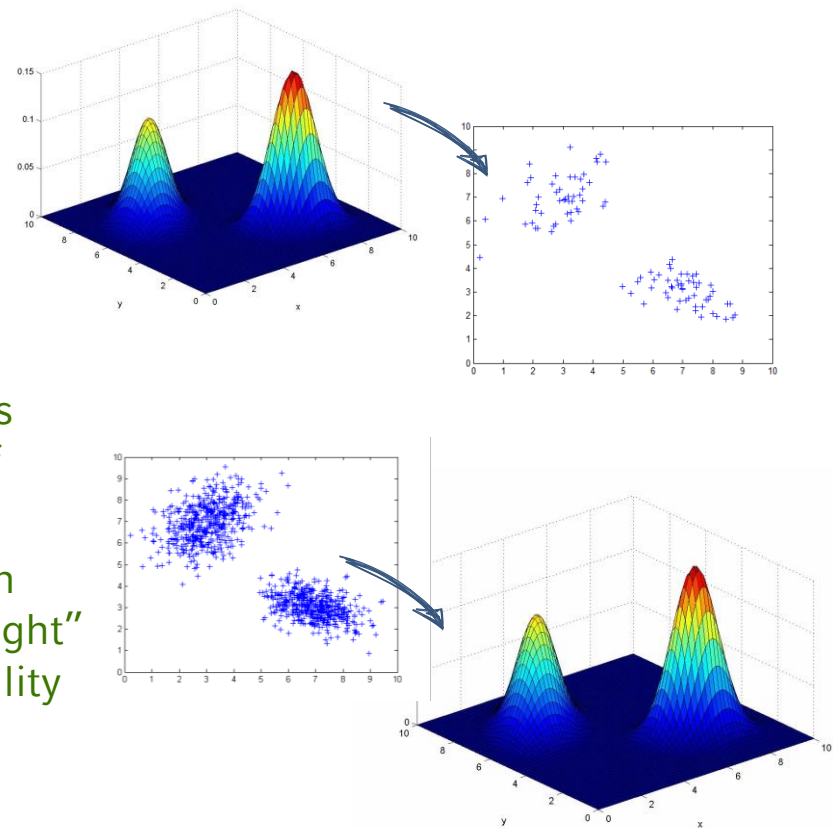
Here: using EM as clustering algorithm

Approach:

Observations are drawn from one of several components of a mixture distribution.

Main idea:

- Define clusters as probability distributions
→ each object has a certain probability of belonging to each cluster
- Iteratively improve the parameters of each distribution (e.g. center, “width” and “height” of a Gaussian distribution) until some quality threshold is reached



Additional Literature: C. M. Bishop „Pattern Recognition and Machine Learning“, Springer, 2009

Excursus: Gaussian Mixture Distributions

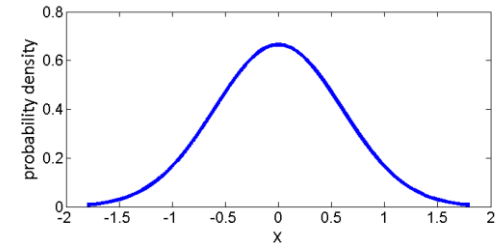
Note: EM is not restricted to Gaussian distributions, but they will serve as example in this lecture.

Gaussian distribution:

- Univariate: single variable $x \in \mathbb{R}$:

$$p(x|\mu, \sigma^2) = \mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{1}{2\sigma^2} \cdot (x-\mu)^2}$$

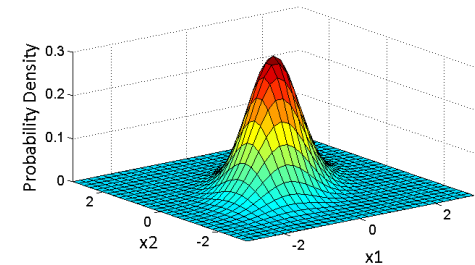
μ \swarrow \nwarrow σ^2
 mean $\in \mathbb{R}$ variance $\in \mathbb{R}$



- Multivariate: d -dimensional vector $\mathbf{x} \in \mathbb{R}^d$:

$$p(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^d |\boldsymbol{\Sigma}|}} \cdot e^{-\frac{1}{2} \cdot (\mathbf{x}-\boldsymbol{\mu})^T \cdot (\boldsymbol{\Sigma})^{-1} \cdot (\mathbf{x}-\boldsymbol{\mu})}$$

$\boldsymbol{\mu}$ \swarrow \nwarrow $\boldsymbol{\Sigma}$
 mean vector $\in \mathbb{R}^d$ covariance matrix $\in \mathbb{R}^{d \times d}$

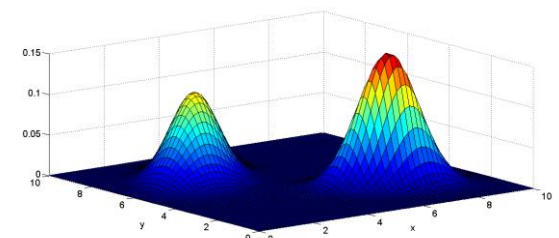


Gaussian mixture distribution with K components:

- d -dimensional vector $\mathbf{x} \in \mathbb{R}^d$:

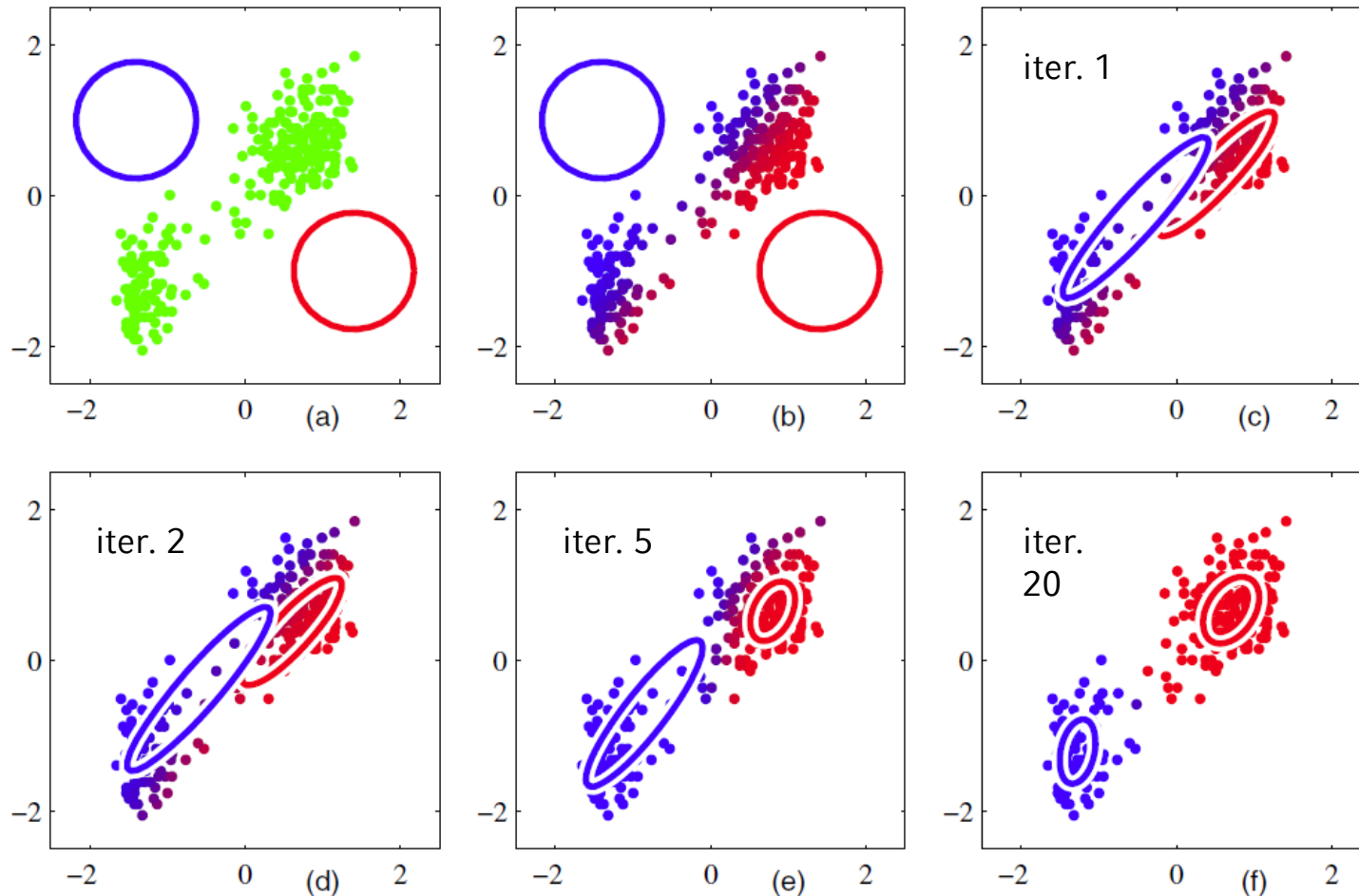
$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k \cdot \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

π_k \uparrow
 mixing coefficients $\in \mathbb{R} : \sum_k \pi_k = 1$ and $0 \leq \pi_k \leq 1$



Expectation Maximization (EM): Exemplary Application

Example taken from: C. M. Bishop „Pattern Recognition and Machine Learning“, 2009



Expectation Maximization (EM)

Note: EM is not restricted to Gaussian distributions, but they will serve as example in this lecture.

A *clustering* $\mathcal{M} = \{C_1, \dots, C_K\}$ is represented by a mixture distribution with parameters $\Theta =$

$$\{\pi_1, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1, \dots, \pi_K, \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K\} :$$

$$p(\mathbf{x}|\Theta) = \sum_{k=1}^K \pi_k \cdot \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

Each *cluster* is represented by one component of the mixture distribution:

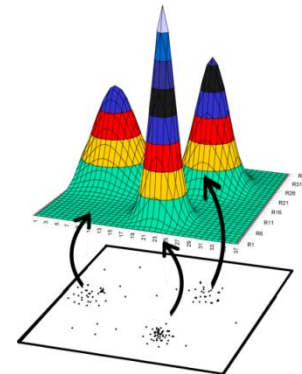
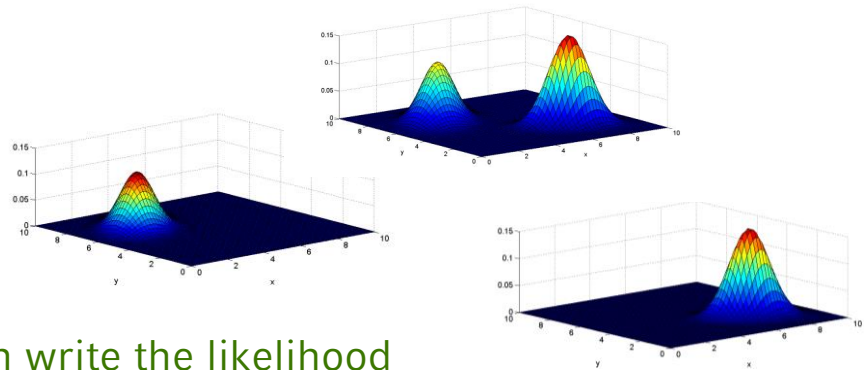
$$p(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

Given a dataset $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\} \subseteq \mathbb{R}^d$, we can write the likelihood that all data points $\mathbf{x}_n \in \mathbf{X}$ are generated (independently) by the mixture model with parameters Θ as:

$$\log p(\mathbf{X}|\Theta) = \log \prod_{n=1}^N p(\mathbf{x}_n|\Theta)$$

Goal: Find the parameters Θ_{ML} with **maximal (log-)likelihood estimation (MLE)**

$$\Theta_{ML} = \arg \max_{\Theta} \{\log p(\mathbf{X}|\Theta)\}$$



Expectation Maximization (EM)

- Goal: Find the parameters Θ_{ML} with the **maximal (log-)likelihood estimation!**

$$\Theta_{ML} = \arg \max_{\Theta} \{\log p(\mathbf{X}|\Theta)\}$$

$$\log p(\mathbf{X}|\Theta) = \log \prod_{n=1}^N \sum_{k=1}^K \pi_k \cdot p(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) = \sum_{n=1}^N \log \sum_{k=1}^K \pi_k \cdot p(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Maximization with respect to the means:

$$\frac{\partial}{\partial \boldsymbol{\mu}_j} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j) = \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_j) \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)$$

$$\frac{\partial \log p(\mathbf{X}|\Theta)}{\partial \boldsymbol{\mu}_j} = \sum_{n=1}^N \frac{\partial \log p(\mathbf{x}_n|\Theta)}{\partial \boldsymbol{\mu}_j} = \sum_{n=1}^N \frac{\frac{\partial p(\mathbf{x}_n|\Theta)}{\partial \boldsymbol{\mu}_j}}{p(\mathbf{x}_n|\Theta)} = \sum_{n=1}^N \frac{\frac{\partial \pi_j \cdot p(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}{\partial \boldsymbol{\mu}_j}}{\sum_{k=1}^K \pi_k \cdot p(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)} = \sum_{n=1}^N \frac{\pi_j \cdot \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_j) \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}{\sum_{k=1}^K \pi_k \cdot p(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

$$\frac{\partial \log p(\mathbf{X}|\Theta)}{\partial \boldsymbol{\mu}_j} = \boldsymbol{\Sigma}_j^{-1} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu}_j) \frac{\pi_j \cdot \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}{\sum_{k=1}^K \pi_k \cdot \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)} \stackrel{\text{def}}{=} \mathbf{0}$$

- Define

$$\gamma_j(\mathbf{x}_n) := \pi_j \cdot \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j).$$

$\gamma_j(\mathbf{x}_n)$ is the probability that component j generated the object \mathbf{x}_n .

Expectation Maximization (EM)

Maximization w.r.t. the means yields:

$$\boldsymbol{\mu}_j = \frac{\sum_{n=1}^N \gamma_j(\mathbf{x}_n) \mathbf{x}_n}{\sum_{n=1}^N \gamma_j(\mathbf{x}_n)} \quad (\text{weighted mean})$$

Maximization w.r.t. the covariance yields:

$$\boldsymbol{\Sigma}_j = \frac{\sum_{n=1}^N \gamma_j(\mathbf{x}_n) (\mathbf{x}_n - \boldsymbol{\mu}_j) (\mathbf{x}_n - \boldsymbol{\mu}_j)^T}{\sum_{n=1}^N \gamma_j(\mathbf{x}_n)}$$

Maximization w.r.t. the mixing coefficients yields:

$$\pi_j = \frac{\sum_{n=1}^N \gamma_j(\mathbf{x}_n)}{\sum_{k=1}^K \sum_{n=1}^N \gamma_k(\mathbf{x}_n)}$$

Problem with finding the optimal parameters Θ_{ML} :

$$\mu_j = \frac{\sum_{n=1}^N \gamma_j(\mathbf{x}_n) \mathbf{x}_n}{\sum_{n=1}^N \gamma_j(\mathbf{x}_n)} \quad \text{and} \quad \gamma_j(\mathbf{x}_n) = \frac{\pi_j \cdot \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)}{\sum_{k=1}^K \pi_k \cdot \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}$$

- Non-linear mutual dependencies.
- Optimizing the Gaussian of cluster j depends on all other Gaussians.
- There is no closed-form solution!
- Approximation through iterative optimization procedures
- Break the mutual dependencies by optimizing μ_j and $\gamma_j(\mathbf{x}_n)$ independently

Expectation Maximization (EM)

EM-approach: iterative optimization

1. Initialize means μ_j , covariances Σ_j , and mixing coefficients π_j and evaluate the initial log likelihood.
2. **E step:** Evaluate the responsibilities using the current parameter values:

$$\gamma_j^{new}(\mathbf{x}_n) = \frac{\pi_j \cdot \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)}{\sum_{k=1}^K \pi_k \cdot \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}$$

3. **M step:** Re-estimate the parameters using the current responsibilities:

$$\mu_j^{new} = \frac{\sum_{n=1}^N \gamma_j^{new}(\mathbf{x}_n) \mathbf{x}_n}{\sum_{n=1}^N \gamma_j^{new}(\mathbf{x}_n)}$$

$$\Sigma_j^{new} = \frac{\sum_{n=1}^N \gamma_j^{new}(\mathbf{x}_n) (\mathbf{x}_n - \mu_j^{new})(\mathbf{x}_n - \mu_j^{new})^T}{\sum_{n=1}^N \gamma_j^{new}(\mathbf{x}_n)}$$

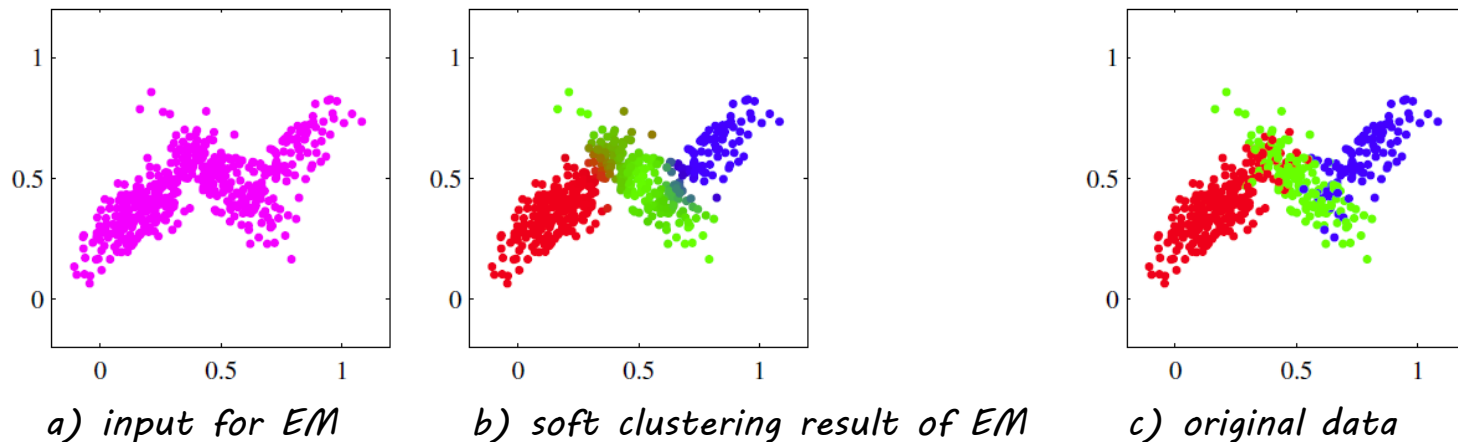
$$\pi_j^{new} = \frac{\sum_{n=1}^N \gamma_j^{new}(\mathbf{x}_n)}{\sum_{k=1}^K \sum_{n=1}^N \gamma_k^{new}(\mathbf{x}_n)}$$

4. Evaluate the new log likelihood $\log p(\mathbf{X} | \Theta^{new})$ and check for convergence of parameters or log likelihood ($|\log p(\mathbf{X} | \Theta^{new}) - \log p(\mathbf{X} | \Theta)| \leq \epsilon$).
If the convergence criterion is not satisfied, set $\Theta = \Theta^{new}$ and go to step 2.

EM: Turning the Soft Clustering into a Partitioning

EM obtains a *soft* clustering (each object belongs to each cluster with a certain probability) reflecting the uncertainty of the most appropriate assignment.

Example taken from: C. M. Bishop „Pattern Recognition and Machine Learning“, 2009



Modification to obtain a *partitioning* variant

- Assign each object to the cluster to which it belongs with the highest probability

$$\text{Cluster}(\text{object}_n) = \underset{k \in \{1, \dots, K\}}{\operatorname{argmax}} \{ \gamma(z_{nk}) \}$$

Superior to k-Means for clusters of varying size
or clusters having differing variances

→ more accurate data representation

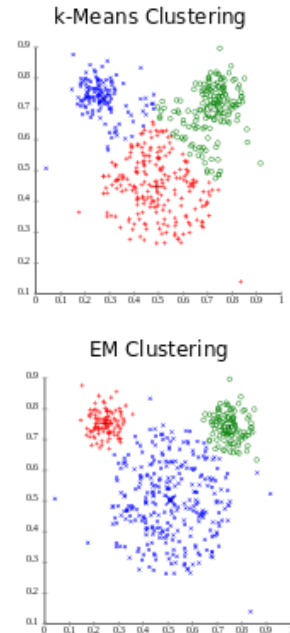
Convergence to (possibly local) maximum

Computational effort for N objects, K derived clusters, and t iterations:

- $O(t \cdot N \cdot K)$
- #iterations is quite high in many cases

Both - result and runtime - strongly depend on

- the initial assignment
 - do multiple random starts and choose the final estimate with highest likelihood
 - Initialize with clustering algorithms (e.g., K-Means usually converges much faster)
 - Local maxima and initialization issues have been addressed in various extensions of EM
- a proper choice of parameter K (= desired number of clusters)
 - Apply principals of model selection (see next slide)



EM: Model Selection for Determining Parameter K

Classical trade-off problem for selecting the proper number of components K

- If K is too high, the mixture may overfit the data
- If K is too low, the mixture may not be flexible enough to approximate the data

Idea: determine candidate models Θ_K for a range of values of K (from K_{min} to K_{max}) and select the model $\Theta_{K^*} = \max\{\text{qual}(\Theta_K) | K \in \{K_{min}, \dots, K_{max}\}\}$

- Silhouette Coefficient (as for k -Means) only works for partitioning approaches.
- The MLE (Maximum Likelihood Estimation) criterion is nondecreasing in K

Solution: deterministic or stochastic *model selection* methods^[MP'00]

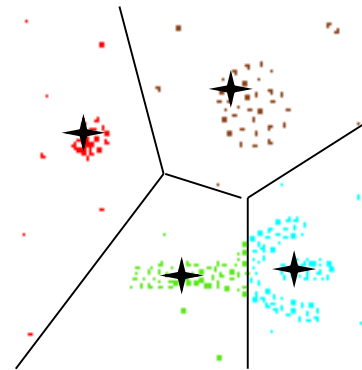
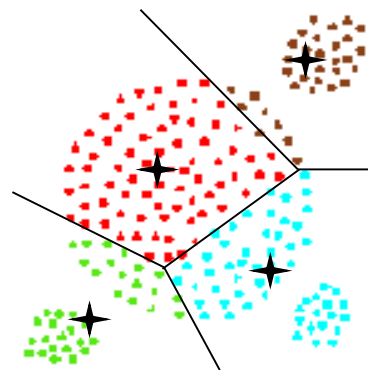
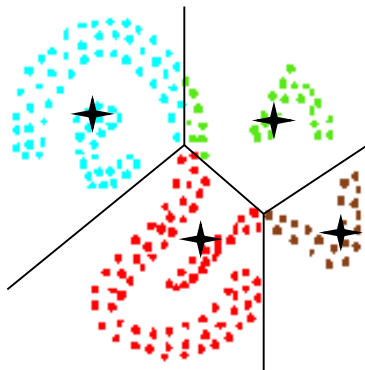
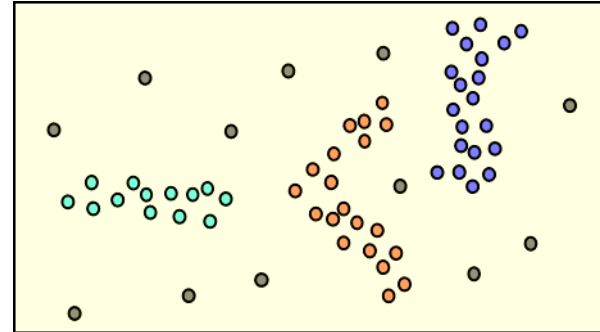
which try to balance the goodness of fit with simplicity.

- Deterministic: $\text{qual}(\Theta_K) = \log p(\mathbf{X} | \Theta_K) + \mathcal{P}(K)$
where $\mathcal{P}(K)$ is an increasing function penalizing higher values of K
- Stochastic: based on Markov Chain Monte Carlo (MCMC)

[MP'00] G. McLachlan and D. Peel. *Finite Mixture Models*. Wiley, New York, 2000.

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- Basic Idea:
 - Clusters are dense regions in the data space, separated by regions of lower object density
- Why Density-Based Clustering?



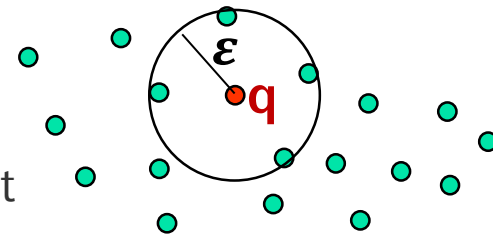
Results of a k -medoid algorithm for $k=4$

- Different density-based approaches exist (see Textbook & Papers)
Here we discuss the ideas underlying the DBSCAN algorithm

Density-Based Clustering: Basic Concept

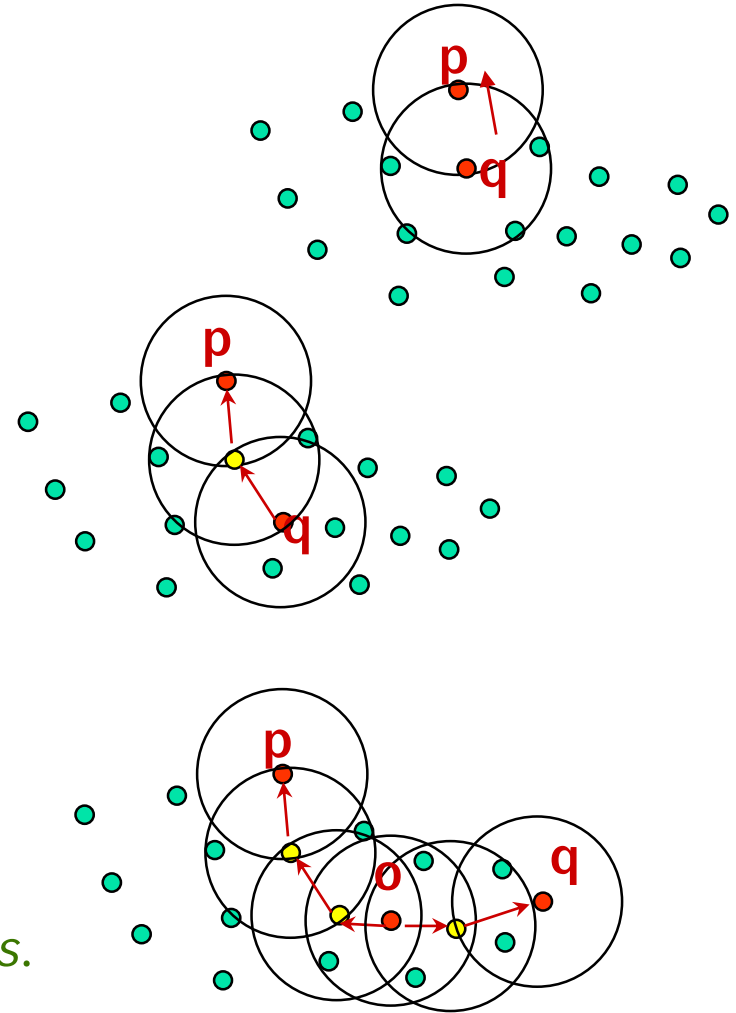
- Intuition for the formalization of the basic idea
 - For any point in a cluster, the local point density around that point has to exceed some threshold
 - The set of points from one cluster is spatially connected
- Local point density at a point q defined by two parameters
 - ε -radius for the neighborhood of point q :
 $N_\varepsilon(q) := \{p \in D \mid \text{dist}(p, q) \leq \varepsilon\}$! contains q itself !
 - **MinPts** – minimum number of points in the given neighbourhood $N_\varepsilon(q)$
- q is called a **core object** (or core point)
 w.r.t. ε , MinPts if $|N_\varepsilon(q)| \geq \text{MinPts}$

$\text{MinPts} = 5 \rightarrow \mathbf{q}$ is a core object



Density-Based Clustering: Basic Definitions

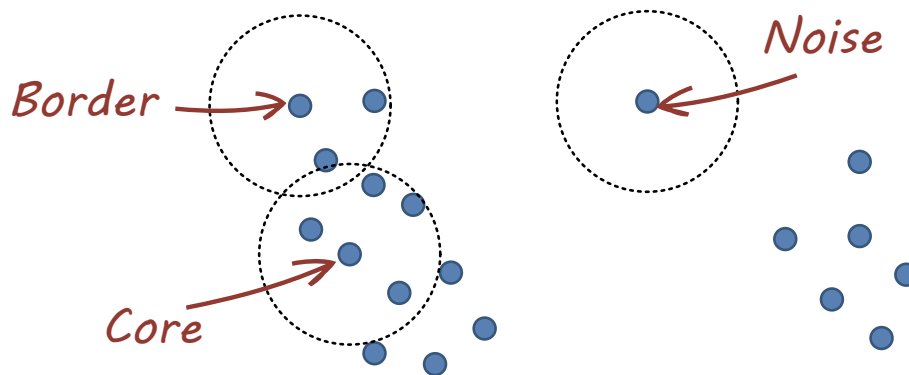
- p **directly density-reachable** from q w.r.t. ε , $MinPts$ if
 - 1) $p \in N_{\varepsilon}(q)$ and
 - 2) q is a core object w.r.t. ε , $MinPts$
- **density-reachable**: transitive closure of *directly* density-reachable
- p is **density-connected** to a point q w.r.t. ε , $MinPts$ if there is a point o such that both, p and q are density-reachable from o w.r.t. ε , $MinPts$.



Density-Based Clustering: Basic Definitions

- **Density-Based Cluster:** non-empty subset S of database D satisfying:
 - 1) *Maximality:* if p is in S and q is density-reachable from p then q is in S
 - 2) *Connectivity:* each object in S is density-connected to all other objects in S

- **Density-Based Clustering** of a database $D : \{S_1, \dots, S_n; N\}$ where
 - S_1, \dots, S_n : all density-based clusters in the database D
 - $N = D \setminus \{S_1 \cup \dots \cup S_n\}$ is called the **noise** (objects not in any cluster)



$\epsilon = 1.0$ $MinPts = 5$

Density-Based Clustering: DBSCAN Algorithm

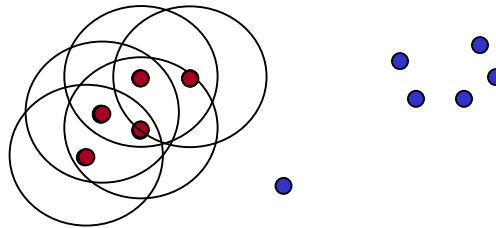
- Density Based Spatial Clustering of Applications with Noise
- Basic Theorem:
 - Each object in a density-based cluster C is density-reachable from any of its core-objects
 - Nothing else is density-reachable from core objects.

```
for each  $o \in D$  do  
    if  $o$  is not yet classified then  
        if  $o$  is a core-object then  
            collect all objects density-reachable from  $o$   
            and assign them to a new cluster.  
        else  
            assign  $o$  to NOISE
```

- density-reachable objects are collected by performing successive ε -neighborhood queries.

Ester M., Kriegel H.-P., Sander J., Xu X.: „A Density-Based Algorithm for Discovering Clusters in Large Spatial Databases with Noise“, In *KDD 1996*, pp. 226—231.

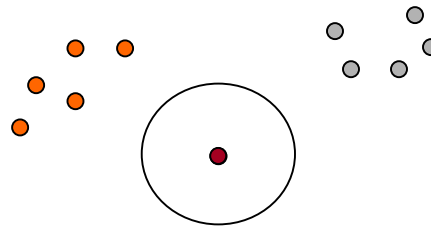
- Parameter
 - $\varepsilon = 2.0$
 - $MinPts = 3$



```

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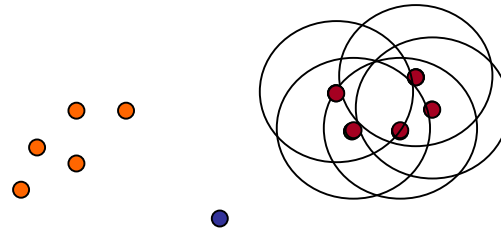
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