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Big Data Analytics Winter Term 2016/2017

Efficient Principal Compenent Analysis (PCA)

Script © 2015 Eirini Ntoutsi, Matthias Schubert, Arthur Zimek

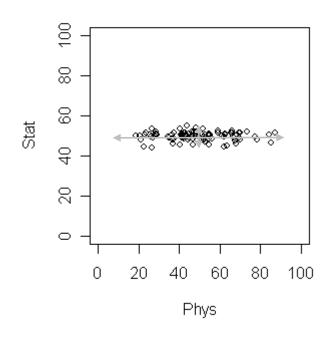
Big Data Management and Analytics



Principal Component Analysis (PCA): A simple example 1/3



- Consider the grades of students in Physics and Statistics.
- If we want to compare among the students, which grade should be more discriminative? Statistics or Physics?



Physics since the variation along that axis is larger.

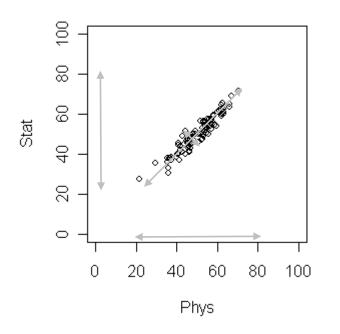
Based on: http://astrostatistics.psu.edu/su09/lecturenotes/pca.html



Principal Component Analysis (PCA): A simple example 2/3



- Suppose now the plot looks as below.
- What is the best way to compare students now?



We should take linear combination of the two grades to get the best results.

Here the direction of maximum variance is clear.

In general ightarrow PCA

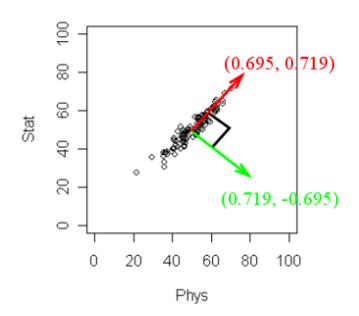
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Principal Component Analysis (PCA): A simple example 3/3



- PCA returns two principal components
 - The first gives the direction of the maximum spread of the data.
 - The second gives the direction of maximum spread perpendicular to the first

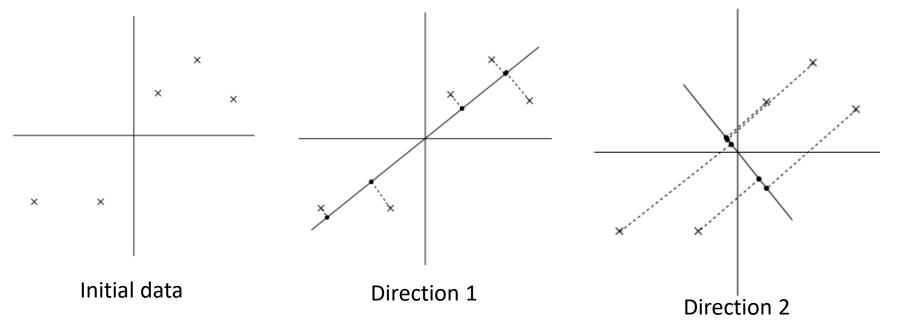


Based on: http://astrostatistics.psu.edu/su09/lecturenotes/pca.html





• The data starts off with some amount of variance/information in it. We would like to choose a direction *u* so that if we were to approximate the data as lying in the direction/subspace corresponding to *u*, as much as possible of this variance is still retained.



Idea: Choose the direction that maximizes the variance of the projected data





- PCA computes the most meaningful basis to re-express a noisy, garbled data set.
- Think of PCA as choosing a new coordinate system for the data, the principal components being the unit vectors along the axes
- PCA asks: Is there another basis, which is a linear combination of the original basis, that best expresses our dataset?
- General form: *PX*=*Y*

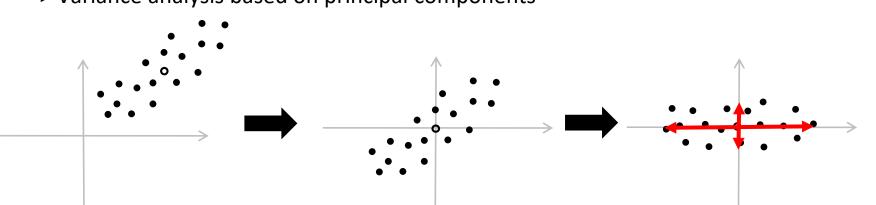
where *P* is a linear transformation, *X* is the original dataset and *Y* the rerepresentation of this dataset.

- P is a matrix that transforms X into Y
- Geometrically, P is a *rotation* and a *stretch* which again transforms X into Y
- The eigenvectors are the rotations to the new axes
- The eigenvalues are the amount of stretching that needs to be done
- The *p*'s are the principal components
 - Directions with the largest variance ... those are the most important, most principal.





Idea: Rotate the data space in a way that the principal components are placed along the main axis of the data space => Variance analysis based on principal components



- Rotate the data space in a way that the direction with the largest variance is placed on an axis of the data space
- Rotation is equivalent to a basis transformation by an orthonormal basis
 - Mapping is equal of angle and preserves distances:

$$x \cdot B = x(b_{*,1}, \dots, b_{*,d}) = (\langle x, b_{*,1} \rangle, \dots, \langle x, b_{*,d} \rangle) \quad mit \quad \forall_{i \neq j} \langle b_i, b_j \rangle = 0 \land \forall_{1 \le i \le d} \|b_i\| = 1$$

• B is built from the largest variant direction which is orthogonal to all previously selected vectors in B.





- Basics of statistical measures:
 - variance
 - covariance
- Basics of linear algebra:
 - Matrices
 - Vector space
 - Basis
 - Eigenvectors, eigenvalues





• A measure of the spread of the data

$$VAR(X) = \frac{1}{n} \sum_{i=1}^{n} (x_i - \mu)^2$$

• Variance refers to a single dimension, e.g., height





• A measure of how much two random variables vary together

$$COV(X,Y) = \frac{1}{n} \sum_{i=1}^{n} (x_i - \mu_x) (y_i - \mu_y)$$

- What the values mean
 - Positive values: both dimensions move together (increase or decrease)
 - Negative values: while one dimension increases the other decreases
 - Zero value: the dimensions are independent of each other.







• Describes the variance of all features and the pairwise correlations between them

$$\Sigma_{D} = \begin{pmatrix} VAR(X_{1}) & \cdots & COV(X_{1}, X_{d}) \\ \vdots & \ddots & \vdots \\ COV(X_{d}, X_{1}) & \cdots & VAR(X_{d}) \end{pmatrix}$$
$$VAR(X) = \frac{1}{n} \sum_{i=1}^{n} (x_{i} - \mu)^{2} \qquad COV(X, Y) = \frac{1}{n} \sum_{i=1}^{n} (x_{i} - \mu_{x}) (y_{i} - \mu_{y})$$

- Properties:
 - For *d*-dimensional data, *dxd* covariance matrix
 - symmetric matrix as COV(X,Y)=COV(Y,X)





• Given *n* vectors $v_i \in IR^d$, $n \times d$ matrix

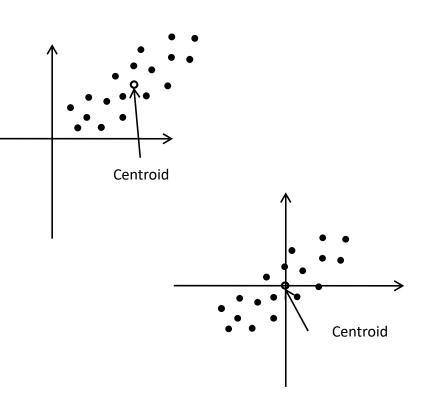
$$D = \begin{pmatrix} v_1 \\ \vdots \\ v_n \end{pmatrix} = \begin{pmatrix} v_{1,1} & \cdots & v_{1,d} \\ \vdots & \ddots & \vdots \\ v_{n,1} & \cdots & v_{n,d} \end{pmatrix}$$

• Centroid/mean vector of D:

$$\vec{\mu} = \frac{1}{n} \cdot \sum_{i=1}^{n} v_i$$

• Centered data matrix:

$$D_{cent} = \begin{pmatrix} v_1 - \vec{\mu} \\ \vdots \\ v_d - \vec{\mu} \end{pmatrix}$$



is called data matrix





• The covariance matrix can be expressed in terms of the centered data matrix as follows:

$$\Sigma_{D} = \begin{pmatrix} VAR(X_{1}) & \cdots & COV(X_{1}, X_{d}) \\ \vdots & \ddots & \vdots \\ COV(X_{d}, X_{1}) & \cdots & VAR(X_{d}) \end{pmatrix} = \frac{1}{n} D_{cent}^{T} D_{cent}$$



Vector/ Matrix basics



• Inner (dot) product of vectors x, y:

$$x \cdot y = x^T \cdot y = (x_1 \quad \cdots \quad x_d) \cdot \begin{pmatrix} y_1 \\ \vdots \\ y_d \end{pmatrix} = \langle x, y \rangle = \sum_{i=1}^d x_i \cdot y_i$$

• Outer product of vectors x, y:

$$\mathbf{x} \otimes \mathbf{y} = \mathbf{x} \cdot \mathbf{y}^{T} = \begin{pmatrix} x_{1} \\ \vdots \\ x_{d} \end{pmatrix} \cdot \begin{pmatrix} y_{1} & \cdots & y_{d} \end{pmatrix} = \begin{pmatrix} x_{1}y_{1} & \cdots & x_{1}y_{d} \\ \vdots & \ddots & \vdots \\ x_{d}y_{1} & \cdots & x_{d}y_{d} \end{pmatrix}$$

• Matrix multiplication:

$$\begin{split} A &= [a_{ij}]_{m \times p}; B = [b_{ij}]_{p \times n}; \\ AB &= C = [c_{ij}]_{m \times n}, where \ c_{ij} = row_i(A) \cdot col_j(B) \end{split}$$

- Length of a vector
 - Unit vector: if ||a||=1

 $\|a\| = \sqrt{a^T \cdot a} = \sqrt{\sum_{i=1}^n a_i^2}$





- Let D be d×d square matrix.
- A non zero vector v_i is called an *eigenvector* of D if and only if there exists a scalar λ_i such that: Dv_i=λ_iv_i.
 - λ_i is called an *eigenvalue* of *D*.
- How to find the eigenvalues/eigenvectors of D?
 - By solving the equation: $det(D-\lambda I_{dxd})=0$ we get the eigenvalues
 - \circ *I_{dxd}* is the identity matrix
 - For each eigenvalue λ_i , we find its eigenvector by solving $(D-\lambda_i)v_i=0$



Eigenvectors decomposition

LMU

- Let D be *dxd* square matrix.
- Eigenvalue decomposition of the data matrix

 $D = V\Lambda V^{T}$ $V = (v_{1}, \dots, v_{d}) \quad mit \forall \forall \langle v_{i}, v_{j} \rangle = 0 \quad und \forall \forall |v_{i}|| = 1$ $\Lambda = \begin{pmatrix} \lambda_{1} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \lambda_{d} \end{pmatrix}$ Every eigenvector is a unit vector The corresponding eigenvalues

- The columns of V are the eigenvectors of D
- The diagonal elements of Λ are the eigenvalues of D





Feature reduction using PCA

- 1. Compute the covariance matrix \varSigma
- 2. Compute the eigenvalues and the corresponding eigenvectors of \varSigma
- 3. Select the *k* biggest eigenvalues and their eigenvectors (V')
- 4. The *k* selected eigenvectors represent an orthogonal basis
- 5. Transform the original $n \times d$ data matrix D with the $d \times k$ basis V':

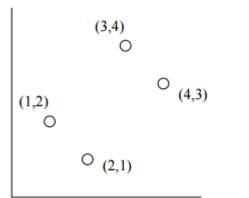
$$D \cdot \mathbf{V}' = \begin{pmatrix} \mathbf{X}_1 \\ \vdots \\ \mathbf{X}_n \end{pmatrix} (v_1', \cdots, v_k') = \begin{pmatrix} \langle \mathbf{X}_1, v_1' \rangle & \cdots & \langle \mathbf{X}_1, v_k' \rangle \\ \vdots & \ddots & \vdots \\ \langle \mathbf{X}_n, v_1' \rangle & \cdots & \langle \mathbf{X}_n, v_k' \rangle \end{pmatrix}$$



Example of transformation



• Original



Eigenvectors $\begin{bmatrix} 1/\sqrt{2} \\ 1/\sqrt{2} \end{bmatrix} \begin{bmatrix} -1/\sqrt{2} \\ 1/\sqrt{2} \end{bmatrix}$

Transformed data

$$\begin{bmatrix} 1 & 2 \\ 2 & 1 \\ 3 & 4 \\ 4 & 3 \end{bmatrix} \begin{bmatrix} 1/\sqrt{2} & -1/\sqrt{2} \\ 1/\sqrt{2} & 1/\sqrt{2} \end{bmatrix} = \begin{bmatrix} 3/\sqrt{2} & 1/\sqrt{2} \\ 3/\sqrt{2} & -1/\sqrt{2} \\ 7/\sqrt{2} & 1/\sqrt{2} \\ 7/\sqrt{2} & 1/\sqrt{2} \end{bmatrix}$$

In the rotated coordinate system

$$(3/\sqrt{2}, 1/\sqrt{2}) \qquad (7/\sqrt{2}, 1/\sqrt{2}) \\ 0 \qquad 0 \\ (3/\sqrt{2}, -1/\sqrt{2}) \qquad (7/\sqrt{2}, -1/\sqrt{2}) \\ (7/\sqrt{2}, -1/\sqrt{2}) \qquad (7/\sqrt{2}, -1/\sqrt{2})$$

Source: http://infolab.stanford.edu/~ullman/mmds/ch11.pdf





- Let k be the number of top eigenvalues out of d (d is the number of dimensions in our dataset)
- The percentage of variance in the dataset explained by the k selected eigenvalues is:

$$\frac{\sum_{i=1}^k \lambda_i}{\sum_{i=1}^d \lambda_i}$$

- Similarly, you can find the variance explained by each principal component
- Rule of thumb: keep enough to explain **85%** of the variation





- Example: iris dataset (d=4), results from R
- 4 principal components

PC1PC2PC3PC4Sepal.Length0.5038236-0.454998720.70885470.19147575Sepal.Width-0.3023682-0.88914419-0.3311628-0.09125405Petal.Length0.5767881-0.03378802-0.2192793-0.78618732Petal.Width0.5674952-0.03545628-0.58290030.58044745

Importance of components: PC1 PC2 PC3 PC4 Proportion of Variance 0.7331 0.2268 0.03325 0.00686 Cumulative Proportion 0.7331 0.9599 0.99314 1.00000

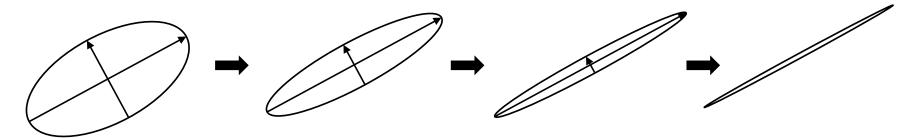




Problem:

- Computing the eigenvalues with standard algorithms is often expensive (many algorithm are well-known)
- Standard methods often in involve matric inversions (O(n³))
- For large matrices more efficient methods are required:
- Most prominent is the power iterations method ($O(n^2)$)

Intuition: Multiplying a matrix with itself increases the strongest direction relative to the other direction.







- given: data n×d matrix X and the corresponding covariance matrix $\Sigma = (X \mu(X)^T(X \mu(X)))$ where $\mu(X)$ is the mean vector of X.
- consider the eigenvalue decomposition of $\Sigma = V^T \Lambda V$ where $V = (v_1, \dots, v_d)$: is the columnwise orthonormal eigenvector basis $\Lambda = \begin{bmatrix} \lambda_1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \lambda_d \end{bmatrix}$: is the diagonal eigenvalue matrix

Note: $\Sigma^t = (V^T \Lambda V)^t = V^T \Lambda V \cdot V^T \Lambda V \cdot \dots \cdot V^T \Lambda V = V^T \Lambda^t V$

$$= V^T \begin{bmatrix} \lambda_1^t & \cdots & 0\\ \vdots & \ddots & \vdots\\ 0 & \cdots & \lambda_d^t \end{bmatrix} \vee$$





- if PCA is well-defined all $\lambda >= 0$
- taking the ith power: All values $\lambda > 1$ increase with the power and all λ values < 0 decrease exponentially fast.
- When normalizing the λ by $\sum_{i=1}^{d} \lambda_i$, we observe the following: for $\lambda_i \neq \lambda_j$ and $t \to \infty$: $\exists \lambda_{i*}: \frac{\lambda_{i*}^t}{\sum_{i=1}^{d} \lambda_i^t} \rightarrow 1$ and $\forall j \neq i *: \frac{\lambda_j^t}{\sum_{i=1}^{d} \lambda_i^t} \rightarrow 0$
- under normalization over all diagonal entries, only one component remains.
- Thus: the rank of Σ^t converges to 1 and the only component remaining is the strongest eigenvector.





The following algorithm computes the strongest eigenvalue of matrix M:

Input: d×d data matrix M x_0 = random unit vector while $x_i / ||x_i|| - |x_{i-1}|| |x_{i-1}|| > \varepsilon$ do $x_i = M^i x_0$ i=i+1return $x_i / ||x_i||$

Why does this work?

$$M^{t}x = \begin{bmatrix} v_{1} & \dots & v_{d} \end{bmatrix} \begin{bmatrix} 0 & \dots & 0 \\ \dots & \lambda_{j}^{t} & \dots \\ 0 & \dots & 0 \end{bmatrix} \begin{bmatrix} v_{1} \\ \vdots \\ v_{d} \end{bmatrix} x = \begin{bmatrix} v_{1} & \dots & v_{d} \end{bmatrix} \begin{bmatrix} 0 & \dots & 0 \\ \dots & \lambda_{j}^{t} & \dots \\ 0 & \dots & 0 \end{bmatrix} \begin{bmatrix} \langle v_{1}, x \rangle \\ \vdots \\ \langle v_{d}, x \rangle \end{bmatrix}$$
$$= \begin{bmatrix} v_{1,1} \cdot 0 + \dots + v_{1,j} \cdot \lambda_{j}^{t} \langle v_{j}, x \rangle + v_{1,d} \cdot 0 \\ \vdots \\ v_{d,1} \cdot 0 + \dots + v_{d,j} \cdot \lambda_{j}^{t} \langle v_{j}, x \rangle + v_{d,d} \cdot 0 \end{bmatrix} = \begin{bmatrix} v_{j} \cdot \lambda_{j}^{t} \langle v_{j}, x \rangle + v_{d,d} \cdot 0 \\ \vdots \\ v_{d,1} \cdot 0 + \dots + v_{d,j} \cdot \lambda_{j}^{t} \langle v_{j}, x \rangle + v_{d,d} \cdot 0 \end{bmatrix}$$

in other words the M^Tx has the same direction as the strongest eigenvector v_i .



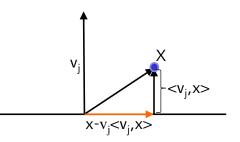


- we now have a method to determine the strongest eigenvalue
- to compute the k-strongest eigenvalues we proceed as follows:
 For i=1 to k:

```
determine the strongest eigenvalue v_i
reproject data X to the space being orthogonal to v_i: x'
= x-v_i<v_i,x>
```

```
output the v_i
```

• explanation for the reprojection:



- if there are two equally strong eigenvalues $\lambda_i = \lambda_j$ then the algorithm return an arbitrary vector from $span(v_i, v_j)$
- for $\lambda_i \approx \lambda_j$: the algorithm converges slower





- PCA is an important method for feature reduction
- general and complete for eigenvalue decomposition are often inefficient(compute the characteristic polynomial, using matrix inversion etc.)
- Power iterations are linear in the size of the matrix, i.e. quadratic in the dimension d.
- Power iterations compute only the k strongest eigenvalues but not all (stop when k strongest v are found)
- rely only on matrix multiplications